Novel pediatric height outlier detection for electronic health records: machine learning with monotonic Bayesian Additive Regression Trees

RA Sparapani, BQ Teng, J Hilbrands, R Pipkorn, MB Feuling, PS Goday Journal of Pediatric Gastroenterology and Nutrition 2002

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Outline

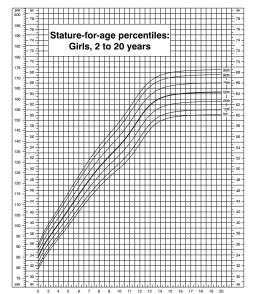
- Motivation: a clinical application in chronically ill children potentially compounded by malnutrition
- ► What is monotonic BART (mBART)
- Nonparametric outlier detection and monotonic advantages
- Nonparametric marginal effect estimation
- Returning to the real data example

Chronically ill children and potential height outliers

- This data is from the electronic health records (EHR) of a large children's health care system
- Chronically ill children are often at high risk for malnutrition
- Typically this is assessed by comparison to Centers for Disease Control (CDC) growth chart benchmarks
- CDC inputs are age, gender, height and weight
- Age and gender are extremely reliable
- However, height and weight are prone to outliers and there is practically NO quality control for these measures i.e., the ground truth of height outliers is largely unknown
- There are about an order of magnitude more height (3%) than weight outliers (0.2%) per measurement (Phan et al. 2020 Scientific Reports)
- Determining malnourishment requires height outlier detection
- Furthermore, this method should be robust to weight outliers that are harder to identify but thankfully less prevalent
- Proposed EHR height outlier removal methods are either too simplistic or too complex to implement (such as Phan et al.)

https://www.cdc.gov/growthcharts

CDC Growth Charts: United States



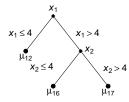
Motivating Example: Growth Chart Outliers

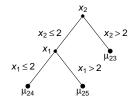
- The US Centers for Disease Control and Prevention (CDC) as well as the World Health Organization (WHO) have developed growth charts for childhood development: height by age, weight by age, body mass index by age and weight by height
- Here we will focus on height, y_t , by age in months, $t = 24, \dots, 215$ (2 to 17 years old)
- The CDC uses the LMS method via natural cubic splines (Cole and Green 1992 Statistics in Medicine)
- Three parameters estimated by penalized maximum likelihood the Box-Cox power transformation, L_t; the mean, M_t; and the coefficient of variation, S_t

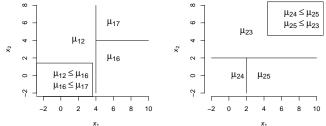
$$z_t = \left\{ \begin{array}{ll} \frac{-1 + (y_t/M_t)^{L_t}}{L_t S_t} & L_t \neq 0\\ \frac{\log(y_t/M_t)}{S_t} & L_t = 0 \end{array} \right\} \sim \mathrm{N}(0, \ 1)$$

- ► CDC/WHO guidelines say values of z_t < -6 or z_t > 6 are outliers but this will catch only the most extreme outliers
- Regardless of the exact cutoff, this outlier method is called Standard Deviation Scores (SDS), i.e., Height SDS

Monotonic example: increasing in x_1 and x_2







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Monotonic BART (mBART)

Chipman et al. 2021 Bayesian Analysis

► $f \stackrel{\text{prior}}{\sim} \text{mBART}$

- A function f is monotone with respect to x_j if f satisfies $f(\ldots, x_{j-1}, x_j + \Delta x, x_{j+1}, \ldots) \ge f(\ldots, x_{j-1}, x_j, x_{j+1}, \ldots)$ for all $\Delta x > 0$ (increasing/nondecreasing) or for all $\Delta x < 0$ (decreasing/nonincreasing)
- Constraint Conditions for Tree Monotonicity
 A tree function g(x; T, M) will be monotone in coordinate x_j
 if the leaf value of each of its terminal node regions is
 (a) not greater than the minimum level of all of its
 above-neighbor regions with respect to x_j and
 (b) not less than the maximum level of all of its
 below-neighbor regions with respect to x_j

Monotonic BART (mBART)

Chipman et al. 2021 Bayesian Analysis

- The leaf prior for BART $\mu_j | \mathcal{T}^{\text{prior}} \sim \mathrm{N}(0, \sigma_{\mu}^2)$
- Consider the simplest case of two monotonic leaves in mBART (relying on the results of Azzalini 1985 Scand J Stat)

$$\begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} \stackrel{\text{prior}}{\sim} N_2 \left(\vec{0}_2, \ c^2 \sigma_{\mu}^2 I_2 \right) \mathbf{I}(\mu_1 < \mu_2) \text{ where } c^2 = \frac{\pi}{\pi - 1} \approx 1.47$$

Equivalent to skew Normal marginals where $V[\mu_1] = V[\mu_2] = \sigma_{\mu}^2$

$$\mu_{1} \stackrel{\text{prior}}{\sim} \phi\left(\frac{\mu_{1}}{c\sigma_{\mu}}\right) \Phi\left(\frac{-\mu_{1}}{c\sigma_{\mu}}\right) \qquad \qquad \mathbf{E}\left[\mu_{1}\right] = \frac{-\sigma_{\mu}}{\sqrt{\pi - 1}}$$
$$\mu_{2} \stackrel{\text{prior}}{\sim} \phi\left(\frac{\mu_{2}}{c\sigma_{\mu}}\right) \Phi\left(\frac{\mu_{2}}{c\sigma_{\mu}}\right) \qquad \qquad \mathbf{E}\left[\mu_{2}\right] = \frac{+\sigma_{\mu}}{\sqrt{\pi - 1}}$$

BART vs. mBART priors

Default BART prior settings
 $\alpha = 0.95, \beta = 2$ I234+Number of leaves1234+Prior probability0.050.550.270.13Default mBART prior settings
 $\alpha = 0.25, \beta = 0.8$ IIIINumber of leavesIIIIIComplete and the state DADET doesIIII

Comparable with BART due to a different sampling approach

Nonparametric outlier detection

- Monotonicity provides additional robustness to outliers since f can't just go up before an outlier and back down after (or vice versa)
- We have *population* predictions of the form
 ŷ_{ij} = E [y_{ij}] = μ + f̂(x_{ij}) where j = 1,..., n_i
 (recall, μ is just a constant roughly centering the population)
- ▶ But these expectations are biased except for the average child
- ▶ We need to adjust these up or down for a given subject
- ► So let m_i = median_j (y_{ij} ŷ_{ij}) (median rather than mean to be robust to outliers)
- ▶ Now, we make *personalized* predictions $\tilde{y}_{ij} = m_i + \hat{y}_{ij}$
- We define the *relative error* of these as $d_{ij} = (y_{ij} \tilde{y}_{ij})/\tilde{y}_{ij}$
- Outliers are defined as |d_{ij}| > δ where δ can be determined from the Receiver Operating Characteristic (ROC) curve
- And the discriminating performance of the method is assessed by the area under the ROC curve

Returning to the real data example

- Constructed two independent cohorts of chronically ill children
 - 2-8 years old
 - measured at least every 120 days on average
 - followed for at least 2 years
- Training cohort: 1376 children with height outliers unknown 39491 measurements: 28.7/child on average
- Validation cohort: 318 children 7378 measurements: 23.2/child on average manually reviewed to determine height outliers however, the ground truth is fallible i.e., retrospective: we can't just re-measure the child's height
- Heights in the Training cohort fit with mBART to age, gender, race/ethnicity and weight

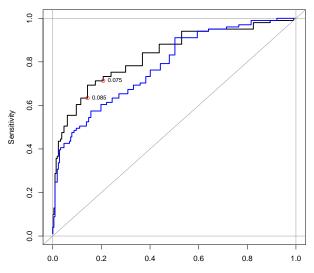
Returning to the real data example

- Outlier detection conducted for the Validation cohort
- The area under the Receiver Operating Characteristic (ROC) curve was excellent: 0.841
- By comparison, if you use the height SDS by age growth chart, the area is only 0.776
- Based on ROC curve, two relative error cutoffs considered Aggressive, 0.075; and Conservative, 0.085

Real data summary

	Training		Validation	
	1376		318	
Children	n	(%)	n	(%)
Female	594	(43.2%)	132	(41.5%)
White	783	(56.9%)	189	(59.4%)
Black	313	(22.7%)	66	(20.8%)
Other	280	(20.3%)	63	(19.8%)
Children with outliers	Unk.		101	(31.8%)
Measurements	m		m	
Height (cm)	39491		7378	
	Mean	(SD)	Mean	(SD)
Measurements/child	28.7		23.2	
First visit age 2	86.4	(8.8)	84.5	(6.6)
Last visit age 5	111.3	(8.4)	109.5	(9.4)
mBART R^2	82.2%		75.3%	

Receiver Operating Characteristic curve (AUC): mBART (0.841) vs. SDS (0.776)



Aggressive cutoff 0.075

- ► B: mBART outlier detection
- ► C: clinical review ground truth

Sensitivity or Recall =
$$P[B = 1|C = 1] = \frac{TP}{Q} = \frac{72}{101} = 0.713$$

Specificity = $P[B = 0|C = 0] = \frac{TN}{M} = \frac{172}{217} = 0.793$
PPV or Precision = $P[C = 1|B = 1] = \frac{TP}{P} = \frac{72}{117} = 0.615$
NPV = $P[C = 0|B = 0] = \frac{TN}{N} = \frac{172}{201} = 0.856$

Conservative cutoff 0.085

- ► B: mBART outlier detection
- ► C: clinical review ground truth

Sensitivity or Recall =
$$P[B = 1|C = 1] = \frac{TP}{Q} = \frac{64}{101} = 0.634$$

Specificity = $P[B = 0|C = 0] = \frac{TN}{M} = \frac{186}{217} = 0.857$
PPV or Precision = $P[C = 1|B = 1] = \frac{TP}{P} = \frac{64}{95} = 0.674$
NPV = $P[C = 0|B = 0] = \frac{TN}{N} = \frac{186}{223} = 0.834$

Aggressive cutoff: targeted smoothing BART

with monotonic weight

Starling et al. Annals of Applied Statistics 2020

- ► B: mBART outlier detection
- ► C: clinical review ground truth

Sensitivity or Recall =
$$P[B = 1|C = 1] = \frac{TP}{Q} = \frac{74}{101} = 0.732$$

Specificity = $P[B = 0|C = 0] = \frac{TN}{M} = \frac{165}{217} = 0.760$
PPV or Precision = $P[C = 1|B = 1] = \frac{TP}{P} = \frac{74}{126} = 0.587$
NPV = $P[C = 0|B = 0] = \frac{TN}{N} = \frac{165}{192} = 0.859$

Aggressive cutoff: females only

- ► B: mBART outlier detection
- ► C: clinical review ground truth

Sensitivity or Recall =
$$P[B = 1|C = 1] = \frac{TP}{Q} = \frac{31}{42} = 0.738$$

Specificity = $P[B = 0|C = 0] = \frac{TN}{M} = \frac{70}{90} = 0.778$
PPV or Precision = $P[C = 1|B = 1] = \frac{TP}{P} = \frac{31}{51} = 0.608$
NPV = $P[C = 0|B = 0] = \frac{TN}{N} = \frac{70}{81} = 0.864$

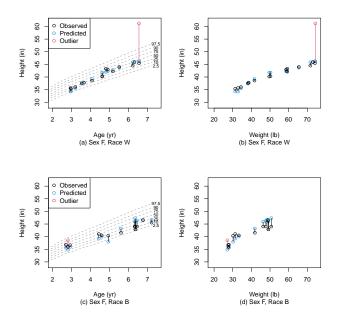
Aggressive cutoff: non-whites only

- ► B: mBART outlier detection
- ► C: clinical review ground truth

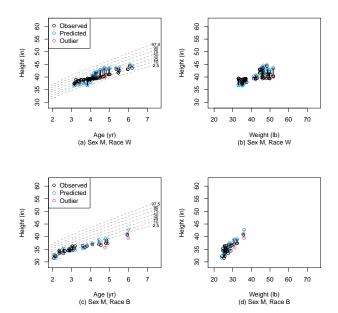
Sensitivity or Recall =
$$P[B = 1|C = 1] = \frac{TP}{Q} = \frac{39}{50} = 0.780$$

Specificity = $P[B = 0|C = 0] = \frac{TN}{M} = \frac{60}{79} = 0.759$
PPV or Precision = $P[C = 1|B = 1] = \frac{TP}{P} = \frac{39}{58} = 0.672$
NPV = $P[C = 0|B = 0] = \frac{TN}{N} = \frac{60}{71} = 0.845$

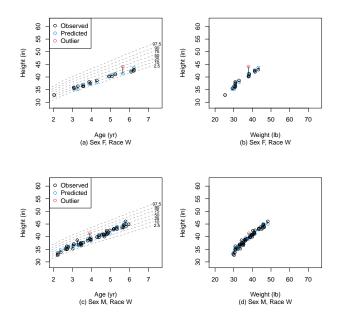
True Positives



False Positives



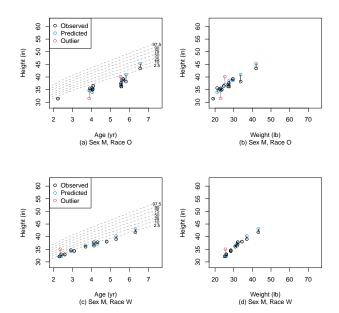
False Negatives



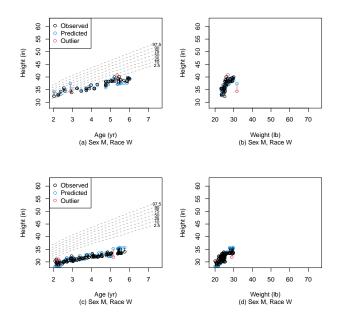
Conclusions

- We constructed our new outlier detection methodology based on nonparametric machine learning via monotonic BART
- This automated method's performance was deemed to be adequate via an indpendent validation cohort
- Modern methodology leads to a simply-tuned single rule as opposed to complex simultaneous tuning of multiple rules that have been proposed based on classic methods
- For EHR heights/weights, the ground truth is unknown prospective corrections are rarely performed and retrospective attempts to identify outliers manually are fallible

True Positives



False Positives



False Negatives

